Regularization Methods for System Identification

Hyperparameter Estimation

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Introduction

Regularization methods have achieved a great success in statistics, machine learning, biometrics, etc, over the last two decades.

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A general framework

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The bias/variance tradeoff is at the heart of identification Suppose that

$$\theta_0$$
 — True parameter $\widehat{\theta}$ — Estimate

Bias-variance tradeoff

$$\underbrace{E \| \widehat{\theta} - \theta_0 \|^2}_{\text{MSE}} = \underbrace{\| E \widehat{\theta} - \theta_0 \|^2}_{\text{bias's square deterministic}} + \underbrace{E \| \widehat{\theta} - E \widehat{\theta} \|^2}_{\text{variance random}}$$

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As complexity of $\mathcal M$ increases, bias decreases but variance increases To choose a proper complexity for the given data and to achieve a "good" bias/variance tradeoff

Linear models

$$Y = \Phi \theta_0 + V$$

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 ℓ_1 -norm regularization

$$\widehat{\theta}_1 \stackrel{\triangle}{=} \arg\min_{\theta \in \mathcal{M}} \left(\| \mathbf{Y} - \mathbf{\Phi} \boldsymbol{\theta} \|^2 + \lambda \| \boldsymbol{\theta} \|_1 \right)$$

To seek parsimonious models: regularization is a prime tool for sparsity

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To seek parsimonious models: regularization is a prime tool for sparsity

 ℓ_2 -norm regularization

$$\widehat{\theta}_2 \stackrel{\triangle}{=} \arg\min_{\theta \in \mathcal{M}} \left(\| \mathbf{Y} - \mathbf{\Phi} \boldsymbol{\theta} \|^2 + \lambda \| \boldsymbol{\theta} \|_2^2 \right)$$

The bias/variance tradeoff is at the heart of identification: regularization offers new techniques for robust smaller MSE

The main issue

Can regularization methods bring forth some benefits for system identification?

Yes!

Regularization methods for linear

system identification

Linear time-invariant (LTI) system identification is a classical and fundamental problem.

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Output error (OE) systems

$$y(t) = \sum_{k=1}^{\infty} g_k^0 u(t-k) + v(t), \ t = 1, 2, \cdots$$

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The Goal

To identify the impulse response sequence

$$\theta_0 = [g_1^0, g_2^0, \cdots]^T$$
 (infinite parameters)

as well as possible by a finite number of data

$$\{u(t), y(t)\}_{t=1}^{N}$$

The impulse response identification could be ill-conditioned in practice since it involves to estimate an infinite number of parameters

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The identification is to make the ill-conditioned problem well-conditioned

Two routes

 Parametric methods (Classical methods: maximum likelihood, prediction error method, etc.)

$$\sum_{k=1}^{\infty} g_k^0 q^{-k} = \frac{b_1 q^{-1} + \dots + b_{n_b} q^{-n_b}}{1 + f_1 q^{-1} + \dots + f_{n_f} q^{-n_f}}$$

- · Model class selection
- · Model order selection: AIC, BIC, cross validation

Asymptotic optimality

· Nonparametric methods

Motivation

 Parametric methods are not as reliable as expected for short, ill-conditioned, low signal-to-noise ratio data

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A high order finite impulse response (FIR) system, (e.g. n = 100)

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A high order finite impulse response (FIR) system, (e.g. n = 100)

$$y(t) = \sum_{k=1}^{n} g_{k}^{0} u(t-k) + v(t)$$

Prior

stability :
$$g_k^0 \sim O(\tau^k)$$
 for some $0 < \tau < 1$

Linear regression form

$$Y = \Phi \theta_0 + V, \ \theta_0 = [g_1^0, g_2^0, \dots, g_n^0]^T$$

where

$$\Phi = \begin{bmatrix} u(0) & u(-1) & \dots & u(-n+1) \\ u(1) & u(0) & \dots & u(-n+2) \\ \vdots & \vdots & \ddots & \vdots \\ u(N-1) & u(N-2) & \dots & u(N-n) \end{bmatrix}$$

$$Y = \begin{bmatrix} y(1) & y(2) & \dots & y(N) \end{bmatrix}^{T}$$

$$V = \begin{bmatrix} v(1) & v(2) & \dots & v(N) \end{bmatrix}^{T}$$

An example

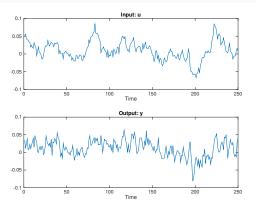
Input-output data of a linear dynamic system:

Data size: 250

· Input: a filtered white noise

Noise: a white noise with the signal to noise ratio 5.45

To estimate the first 100 impulse response coefficients



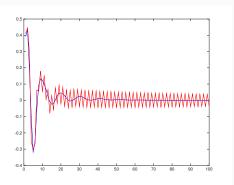
Performance measure

$$Fit = 100 \times \left(1 - \frac{\|\widehat{\theta}_{im} - \theta_0\|}{\|\theta_0 - \bar{\theta}_0\|}\right), \ \bar{\theta}_0 = \frac{1}{n} \sum_{k=1}^n g_k^0$$

where $\hat{\theta}_{im}$ is the corresponding first n=100 impulse response of the estimate for $\hat{\theta}$.

The OE-system of order 6 by CV

Fit = 36.78

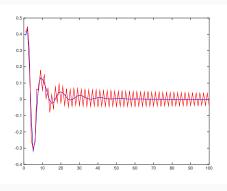


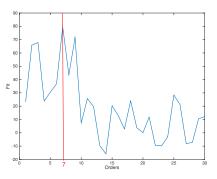
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The best OE system of the order 7

Fit = 79.63



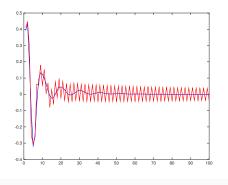


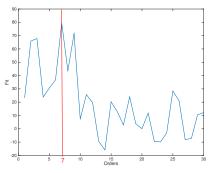
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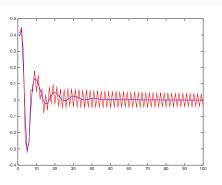
The estimate is sensitive to the choice of model order

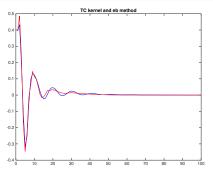
The OE-system of order 6 by CV

Fit = 36.78

Regularization methods

Fit = 83.40





Objective functions

$$\underbrace{\ell(Y,\Phi\theta)}_{\text{loss term}} + \underbrace{R(\theta)}_{\text{regularization term}}$$

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Loss term

· characterize the feature of the noise

Regularization term

- ill-posed problem
- encode prior knowledge

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Loss term

· characterize the feature of the noise

Regularization term

- · ill-posed problem
- encode prior knowledge

Some examples

$$\|Y - \Phi\theta\|_p^p + \lambda \|\theta\|_q^q, \ p \ge 0, q \ge 0$$

Recall that

$$\widehat{\boldsymbol{\theta}} \stackrel{\triangle}{=} \arg \min_{\boldsymbol{\theta} \in \mathcal{M}} \left(\text{Fit + Complexity penalty} \right)$$

Linear regression

$$Y = \Phi \theta_0 + V, \quad \theta_0 = [g_1^0, g_2^0, \dots, g_n^0]^T$$
$$y(t) = \sum_{k=1}^n g_k^0 u(t-k) + v(t)$$

Least squares (LS) estimators:

$$\begin{split} \widehat{\theta}^{\mathrm{LS}} &\stackrel{\triangle}{=} \arg\min_{\theta} \| Y - \Phi^T \theta \|^2 = (\Phi^T \Phi)^{-1} \Phi^T Y \\ \mathsf{MSE}(\widehat{\theta}^{\mathrm{LS}}) &= E \| \widehat{\theta}^{\mathrm{LS}} - \theta_0 \|^2 = \sigma^2 \mathsf{Tr} \big((\Phi^T \Phi)^{-1} \big) \end{split}$$

Too many parameters? Put them on leashes!

$$\widehat{\theta}^{\mathrm{R}} \stackrel{\triangle}{=} \arg \min_{\theta \in \mathbb{R}^n} \| Y - \Phi \theta \|^2 + \frac{\sigma^2 \theta^\mathsf{T} \mathsf{K}^{-1} \theta}{} = (\Phi^\mathsf{T} \Phi + \frac{\sigma^2 \mathsf{K}^{-1}}{})^{-1} \Phi^\mathsf{T} Y$$

where K is a positive semidefinite matrix to be tuned by the data.

A frequentist perspective

The estimator:

$$\widehat{\theta}^{\mathrm{R}} \stackrel{\triangle}{=} \arg \min_{\theta \in \mathbb{R}^n} \| Y - \Phi \theta \|^2 + \sigma^2 \theta^\mathsf{T} K^{-1} \theta = R^{-1} \Phi^\mathsf{T} Y, \ R = \Phi^\mathsf{T} \Phi + \sigma^2 K^{-1}$$

Bias

$$E\widehat{\theta}^{\mathrm{R}} - \theta_0 = \sigma^2 R^{-1} K^{-1} \theta_0 \neq 0$$

MSE

$$E\|\widehat{\theta}^{R} - \theta_0\|^2 = \underbrace{\sigma^4 \theta_0^T P^{-1} R^{-1} R^{-1} K^{-1} \theta_0}_{\text{bias's square}} + \underbrace{\sigma^2 \text{Tr} \left(R^{-1} \Phi^T \Phi R^{-1}\right)}_{\text{variance}}$$

No regularization if $K^{-1}=0$: Bias = 0 and Variance = $\sigma^2(\Phi^T\Phi)^{-1}$ **Proposition**

If $\sigma^2 K^{-1} = \beta A$ and A is positive definite and fixed. Then we have

$$MSE(\widehat{\theta}^{R}) \leq MSE(\widehat{\theta}^{LS}), \text{ when } 0 < \beta < 2\sigma^{2}/(\theta_{0}^{T}A\theta_{0})$$

The optimal kernel matrix for any data length

$$K = \theta_0 \theta_0^T$$

Bayesian Interpretation

Prior

$$\theta_0 \sim \mathcal{N}(0, K)$$
 (K : Covariance/Kernel matrix)

Posterior

$$\begin{aligned} &\theta_0 | Y \sim \mathcal{N}(\widehat{\theta}^{R}, \widehat{K}^{R}) \\ &\widehat{\theta}^{R} = R^{-1} \Phi^T Y, \ \widehat{K}^{R} = \sigma^2 R^{-1} \\ &R = \Phi^T \Phi + \frac{\sigma^2 K^{-1}}{2} \end{aligned}$$

This interpretation provides a clue to select K

Regularization in system identification

Regularization for handling ill-posed problems (Tikhonov & Arsenic, 1977)¹

¹A. N. Tikhonov and V. Y. Arsenic. Solutions of Ill-Posed Problems, New York: John Wiley, 1977.

²J. Sjöberg, T. McKelvey, and L. Ljung. On the use of regularization in system identification. Proceedings of the 12th IFAC World Congress: 381–386, Sydney, Australia.

³G. Pillonetto and G. De Nicolao. A new kernel-based approach for linear system identification. *Automatica*, 46, 81–93, 2010.

Regularization in system identification

Regularization for handling ill-posed problems (Tikhonov & Arsenic, 1977)¹

Regularization is not new in system identification

The first paper in system identification (Sjöberg et al., 1993)²

$$\begin{split} \widehat{\theta}^{R} &= \arg\min_{\theta} \|Y - \Phi \theta\|^{2} + \gamma \|\theta\|^{2} \\ &= (\Phi^{T} \Phi + \gamma I_{n})^{-1} \Phi^{T} Y \end{split}$$

But no important progress until Pillonetto & De Nicolao (2010)³

¹A. N. Tikhonov and V. Y. Arsenic. Solutions of Ill-Posed Problems, New York: John Wiley, 1977.

²J. Sjöberg, T. McKelvey, and L. Ljung. On the use of regularization in system identification. Proceedings of the 12th IFAC World Congress: 381–386, Sydney, Australia.

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How to tune a "good" kernel K by the data

The estimator:

$$\widehat{\theta}^{\mathrm{R}} = (\Phi^{\mathsf{T}}\Phi + \sigma^{2}K^{-1})^{-1}\Phi^{\mathsf{T}}Y$$

Two extrema

$$\widehat{\theta}^{\mathrm{R}} = \left\{ \begin{array}{ll} 0, & \text{if } K = 0 \\ \widehat{\theta}^{\mathrm{LS}}, & \text{if } K = \infty \end{array} \right.$$

How to tune a "good" kernel K by the data

A two-step procedure

The seminal paper (Pillonetto & De Nicolao, 2010)¹

 Kernel design: determine the structure of K by using the prior knowledge

$$K(\eta)$$
, η hyperparameter

 Hyperparameter estimation: determine the hyperparameter by the data

¹G. Pillonetto and G. De Nicolao. A new kernel-based approach for linear system identification. *Automatica*, 46, 81–93, 2010.

Kernel design

Cubic spline kernels (Wahba, 1990)¹

$$K_{CS}(i,j) = \begin{cases} c\frac{i^2}{2} \left(j - \frac{i}{3} \right), & i \ge j \\ c\frac{j^2}{2} \left(i - \frac{j}{3} \right), & i < j \end{cases}$$

Prior: exponential decay

$$g_k^0 \sim O(au^k)$$
 for some $0 < au < 1$

Stable spline kernels (Pillonetto & De Nicolao, 2010)²

An exponential transform:

$$i \to \lambda^i$$

for some $0 < \lambda < 1$

$$K_{SS}(i,j) = \begin{cases} c \frac{\lambda^{2i}}{2} \left(\lambda^{j} - \frac{\lambda^{i}}{3} \right), & i \geq j \\ c \frac{\lambda^{2j}}{2} \left(\lambda^{i} - \frac{\lambda^{j}}{3} \right), & i < j \end{cases}$$

¹G. Wahba. Spline Models for Observational Data. New York: SIAM, 1990.

²G. Pillonetto and G. De Nicolao. A new kernel-based approach for linear system identification. *Automatica*, 46, 81–93, 2010.

The optimal kernel

$$K = \theta_0 \theta_0^{\mathsf{T}} = \begin{bmatrix} (g_1^0)^2 & g_1^0 g_2^0 & \cdots & g_1^0 g_n^0 \\ g_2^0 g_1^0 & (g_2^0)^2 & \cdots & g_2^0 g_n^0 \\ \vdots & \ddots & \ddots & \vdots \\ g_n^0 g_1^0 & g_n^0 g_2^0 & \cdots & (g_n^0)^2 \end{bmatrix}$$

$$\theta_0 = [g_1^0, \cdots, g_n^0]^{\mathsf{T}}$$

Prior

$$g_k^0 \sim O(\tau^k)$$
 for some $0 < \tau < 1$

DI kernel

$$K(\eta) = c \operatorname{diag}([\lambda, \dots, \lambda^n])$$
$$\eta = [c, \lambda] \in \Omega = \{c \ge 0, 0 \le \lambda \le 1\}$$

DI kernel

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$$\eta = [c, \lambda] \in \Omega = \{c \ge 0, 0 \le \lambda \le 1\}$$

DC kernel

$$K_{i,j}(\eta) = c \lambda^{(i+j)/2} \rho^{|i-j|}$$

$$K(\eta) = c \begin{bmatrix} \lambda & \lambda^{\frac{3}{2}} \rho & \cdots & \lambda^{\frac{n+1}{2}} \rho^{n-1} \\ \lambda^{\frac{3}{2}} \rho & \lambda^2 & \cdots & \lambda^{\frac{n+2}{2}} \rho^{n-2} \\ \vdots & \ddots & \ddots & \vdots \\ \lambda^{\frac{n+1}{2}} \rho^{n-1} & \lambda^{\frac{n+2}{2}} \rho^{n-2} & \cdots & \lambda^n \end{bmatrix}$$

with hyperparameters $\eta = [c, \lambda, \rho]^T \in \Omega = \{c \ge 0, 0 \le \lambda \le 1, |\rho| \le 1\}.$

TC kernel (Chen et al., 2012) 1

A special case of DC kernel with $\rho = \sqrt{\lambda}$.

$$K_{k,j}(\eta) = c \min(\lambda^k, \lambda^j), K(\eta) = c \begin{bmatrix} \lambda & \lambda^2 & \cdots & \lambda^n \\ \lambda^2 & \lambda^2 & \cdots & \lambda^n \\ \vdots & \ddots & \ddots & \vdots \\ \lambda^n & \lambda^n & \cdots & \lambda^n \end{bmatrix}$$

with hyperparameters $\eta = [c, \lambda]^T \in \Omega = \{c \ge 0, 0 \le \lambda \le 1\}.$

¹T. Chen, H. Ohlsson, and L. Ljung. On the estimation of transfer functions, regularizations and Gaussian processes–Revisited. *Automatica*, 48(8): 1525–1535, 2012.

Multiple kernels (Chen et al., 2014)1

Better capture complicated dynamics of the system

$$K(\eta) = \sum_{i=1}^{m} \mathbf{\eta}_i K_i, \ \eta = [\eta_1, \cdots, \eta_m]$$

where K_i has different dynamic behavior, e.g. decaying rate and magnitude.

¹T. Chen, M. S. Andersen, L. Ljung, A. Chiuso, and G. Pillonetto. System identification via sparse multiple kernel-based regularization using sequential convex optimization techniques. *IEEE Transactions on Automatic Control*, *59*(11): 2933–2945, 2014.

Hyperparameter estimation

The goal

· To estimate the hyperparameters based on the data

The essence

• To tune model complexity in a continuous way

Some commonly used methods (Pillonetto et al., 2014) ¹

- 1. Empirical Bayes (EB)
- 2. Stein's unbiased risk estimator (SURE)
- 3. Cross validation (CV)

¹G. Pillonetto, F. Dinuzzo, T. Chen, G. De Nicolao, and L. Ljung. Kernel methods in system identification, machine learning and function estimation: A survey. *Automatica*, *50*(3): 657–682, 2014.

Empirical Bayes

Gaussian prior

$$\theta \sim \mathcal{N}(0, K)$$

$$Y = \Phi\theta + V \sim \mathcal{N}(0, Q)$$

$$Q = \Phi K \Phi^{T} + \sigma^{2} I_{N}$$

Empirical Bayes (EB)

$$\mathrm{EB}: \widehat{\eta}_{\mathrm{EB}} = \arg\min_{\eta \in \Omega} \mathbf{Y}^{\mathsf{T}} \mathbf{Q}^{-1} \mathbf{Y} + \log \det(\mathbf{Q})$$

Stein's unbiased risk estimator (SURE)

MSE (for prediction ability):

$$MSE(K) = E \| \mathbf{\Phi}(\widehat{\theta}^{R} - \theta_{0}) \|^{2}$$

It is intractable to tune the hyperparameter by the MSE in practice

SURE method

To construct an unbiased estimators of the MSE

$$\mathcal{F}_{SURE}(K) = \|Y - \Phi \widehat{\theta}^{R}\|^{2} + \frac{2}{3}\sigma^{2} \operatorname{Tr}(R^{-1}\Phi^{T}\Phi)$$

$$R = \Phi^{T}\Phi + \sigma^{2}K^{-1}$$

• To estimate the hyperparameter η by

$$\mathsf{SURE}: \ \widehat{\eta}_{\mathsf{SURE}} = \arg\min_{\eta \in \Omega} \mathscr{F}_{\mathsf{SURE}}(\mathit{K}(\eta))$$

Cross-validation

Ideas

- · divide the whole data into training data and validation data
- · estimate on the training data
- · evaluate on the validation data

Averaged prediction error

For each splitting way s,

- · s the index set of the validation data
- \cdot s^c the index set of the training data

where

$$|s| = k, \{1, \dots, N\} = s \cup s^{c}$$

the averaged prediction error (APE) over the validation data is

$$APE_{s} = \frac{1}{k} \sum_{t \in s} (y(t) - \phi(t)^{\mathsf{T}} \widehat{\theta}_{s^{c}})^{2} = \frac{1}{k} ||Y_{s} - \Phi_{s} \widehat{\theta}_{s^{c}}||^{2}$$

Advantage: does not require to estimate the noise variance σ^2

Variants of CVs

1. Leave-k-out cross validation (LKOCV, intractable in general)

$$\widehat{\eta}_{\text{LKOCV}} = \arg\min_{\eta \in \Omega} \frac{1}{\binom{N}{k}} \sum_{\text{S}} \text{APE}_{\text{S}} \text{ (all choices)}$$

2. Leave-one-out cross validation (LOOCV) (k = 1)

$$\widehat{\eta}_{\text{LOOCV}} = \arg\min_{\eta \in \Omega} \frac{1}{N} \sum_{\text{S}} \text{APE}_{\text{S}} = \arg\min_{\eta \in \Omega} \frac{1}{N} \sum_{t=1}^{N} \left(\frac{y(t) - \widehat{y}(t)}{1 - h_{tt}} \right)^2$$

where $H = \Phi(\Phi^T \Phi + \sigma^2 K^{-1})^{-1} \Phi^T$.

3. Generalized cross validation (GCV)

$$\widehat{\eta}_{\text{GCV}} = \arg\min_{\eta \in \Omega} \frac{1}{N} \frac{\sum_{t=1}^{N} (y(t) - \widehat{y}(t))^{2}}{\left(1 - \text{Tr}(H)/N\right)^{2}}$$

The key issue

How to choose a proper hyperparameter estimator for a given data?

Asymptotically theoretical properties

Suppose that

$$\Phi^T \Phi/N \to \Sigma > 0$$
 as $N \to \infty$.

Then the asymptotically optimal hyperparameter in the MSE sense is (Mu et al., 2018c) ¹

$$\boldsymbol{\eta^*} = \arg\min_{\boldsymbol{\eta} \in \Omega} \boldsymbol{\theta}_0^T \boldsymbol{K}^{-1} \boldsymbol{\Sigma}^{-1} \boldsymbol{K}^{-1} \boldsymbol{\theta}_0 - 2 \mathrm{Tr} \big(\boldsymbol{\Sigma}^{-1} \boldsymbol{K}^{-1} \big)$$

depending on the true parameter, chosen kernel, and asymptotic covariance of the input

¹B. Mu, T. Chen and L. Ljung. On Asymptotic Properties of Hyperparameter Estimators for Kernel-based Regularization Methods. Automatica, 94: 381–395, 2018.

Asymptotically theoretical properties

Theorem

- $\begin{array}{l} \boldsymbol{\cdot} \ \widehat{\eta}_{\text{SURE}} \rightarrow \boldsymbol{\eta}^* \\ \widehat{\eta}_{\text{EB}} \rightarrow \arg\min_{\boldsymbol{\eta} \in \Omega} \boldsymbol{\theta}_0^\mathsf{T} \boldsymbol{K}^{-1} \boldsymbol{\theta}_0 + \log \det(\boldsymbol{K}) \ \ (\text{Mu et al., 2018c})^{\; 1} \end{array}$
- $\widehat{\eta}_{\text{GCV}} \to \eta^*$ (Mu et al., 2018a) ² $\widehat{\eta}_{\text{LOOCV}} \to \eta^*$ if the input is bounded $\widehat{\eta}_{\text{LKOCV}} \to \eta^*$ if $k/N \to 0$ and the input is bounded (Mu et al., 2018b) ³

¹B. Mu, T. Chen and L. Ljung. On Asymptotic Properties of Hyperparameter Estimators for Kernel-based Regularization Methods. Automatica, 94: 381–395, 2018.

²B. Mu, T. Chen and L. Ljung. Asymptotic Properties of Generalized Cross Validation Estimators for Regularized System Identification. Proceedings of the IFAC Symposium on System Identification, 203–205, 2018.

³B. Mu, T. Chen and L. Ljung. Asymptotic Properties of Hyperparameter Estimators by Using Cross-Validations for Regularized System Identification. Proceedings of the IEEE Conference on Decision and Control, 644–649, 2018.

Numerical illustrations

Systems: 1000 30th order OE test systems

3 Inputs:

- · IT1, white Gaussian noise
- IT2, white Gaussian noise filtered by $1/(1 0.95q^{-1})^2$
- IT3, the impulsive input, $[\sqrt{N}, 0, \dots, 0]$ (unbounded)

Noises: The SNR is uniformly distributed over [1, 10]

Sample sizes: N = 500,8000

Kernel: TC kernel

Tuning methods:

- · EB, LOOCV, GCV, SURE
- MSE for reference (optimal for any finite sample)

Results

Table 1: Average fits for 1000 test systems.

Inputs	Sizes	EB	LOOCV	GCV	SURE	MSE
IT1	500	86.16	86.24	86.24	86.03	87.02
	8000	96.44	96.60	96.60	96.60	96.67
IT2	500	39.03	-85.95	-84.84	-146.4	41.94
	8000	50.86	38.79	38.89	38.86	53.63
IT3	500		69.33	89.55	89.52	89.95
	8000		81.42	96.64	96.64	96.70

Conclusion

Summary

- A brief introduction of regularization methods for impulse response identification of linear dynamic systems is given.
- Asymptotically theoretical properties of several hyperparameter estimation are shown.

Thanks for your listening



References

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